

Asymptotics of Mahonian statistics

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Slides: http://www.math.ucsd.edu/~jswanson/talks/2020_UCSD.pdf

Inversion number definition

Definition

The *symmetric group* is

$$S_n := \{\text{bijections } \pi: \{1, 2, \dots, n\} \rightarrow \{1, 2, \dots, n\}\}.$$

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$$\text{inv}(\pi) := \#\{(i, j) : 1 \leq i < j \leq n, \pi(i) > \pi(j)\}.$$

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For example, using *one-line notation* $\pi(1)\pi(2) \cdots \pi(n)$,

$$\text{inv}(35142) = \#\{(1, 3), (1, 5), (2, 3), (2, 4), (2, 5), (4, 5)\} = 6.$$

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Zeilberger: this is the “**most important permutation statistic**”.

Inversion number bijection

Lemma (Classical)

The map

$$\Phi: S_n \rightarrow \{\alpha \in \mathbb{Z}_{\geq 0}^n : (\alpha_1, \dots, \alpha_n) \leq (n-1, n-2, \dots, 1, 0)\}$$
$$\alpha_i := \#\{j : i < j \leq n, \pi(i) > \pi(j)\}$$

*is a **bijection**.*

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	(2,5)		
(1,5)	(2,4)		
(1,3)	(2,3)		(4,5)

$$\Rightarrow \Phi(35142) = (2, 3, 0, 1, 0)$$

Inversion number distribution

Clearly $\text{inv}(\pi) = \alpha_1 + \cdots + \alpha_n$. Hence:

Corollary

- ▶ $\mathcal{X}_{\text{inv}} \sim \mathcal{U}_{n-1} + \cdots + \mathcal{U}_1$ is the **sum of independent discrete uniform random variables**.

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- ▶ inv on S_n is **symmetrically distributed** with mean $((n-1) + (n-2) + \cdots + 0)/2 = \frac{1}{2} \binom{n}{2}$.

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- ▶ inv on S_n is **symmetrically distributed** with mean $((n-1) + (n-2) + \cdots + 0)/2 = \frac{1}{2} \binom{n}{2}$.
- ▶ The ordinary generating function of inv on S_n is

$$\begin{aligned} \sum_{\pi \in S_n} q^{\text{inv}(\pi)} &= \prod_{i=1}^n \sum_{\alpha_i=0}^{n-i} q^{\alpha_i} \\ &= (1 + q + \cdots + q^{n-1})(1 + q + \cdots + q^{n-2}) \cdots (1) \\ &=: [n]_q! \end{aligned}$$

Inversion number distribution

Theorem (Feller '45; implicit earlier)

As $n \rightarrow \infty$, \mathcal{X}_{inv} is *asymptotically normal*.

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As $n \rightarrow \infty$, \mathcal{X}_{inv} is *asymptotically normal*. That is, for all $u \in \mathbb{R}$,

$$\lim_{n \rightarrow \infty} \mathbb{P}[\mathcal{X}_{\text{inv}}^* \leq u] = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^u e^{-x^2/2} dx$$

where

$$\mathcal{X}_{\text{inv}}^* := \frac{\mathcal{X}_{\text{inv}} - \mu_n}{\sigma_n}$$

with

$$\mu_n = \frac{n(n-1)}{4}, \quad \sigma_n^2 = \frac{2n^3 + 3n^2 - 5n}{72}.$$

Inversion number application

Theorem (Feller '45; implicit earlier)

As $n \rightarrow \infty$, \mathcal{X}_{inv} is *asymptotically normal*.

Example (*Kendall's τ test*)

Say some process generated distinct real numbers x_1, x_2, \dots, x_n , one per day. You want to know if the process is **independent of time**. Turn the data into a permutation π while preserving the relative order of data points and compute $\text{inv}(\pi)$. Since $\mathcal{X}_{\text{inv}} \approx \mathcal{N}(\mu_n, \sigma_n)$, independent data would have

$$|(\text{inv}(\pi) - \mu_n)/\sigma_n| \leq 3$$

$\approx 99.7\%$ of the time. So, **if this z-score is too big**, say larger than 3, the process is very likely time-dependent.

Major index definition

Definition (MacMahon, early 1900's)

The *descent set* of $\pi \in S_n$ is

$$\text{Des}(\pi) := \{1 \leq i \leq n-1 : \pi(i) > \pi(i+1)\}.$$

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$$\text{maj}(25143) = \text{maj}(25.14.3) = 2 + 4 = 6.$$

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For example,

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Zeilberger: this is the “**second most important permutation statistic**”.

Major index bijection

Lemma (Gupta, '78)

For a given $\pi \in S_{n-1}$, let $C_\pi \subset S_n$ be the n permutations obtained by inserting n into π in all possible ways. Then

$$\{\text{maj}(\pi') - \text{maj}(\pi) : \pi' \in C_\pi\} = \{0, 1, \dots, n-1\}.$$

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$$\text{maj}(21453) = 1 + 4 = 5$$

$$\text{maj}(21435) = 1 + 3 = 4.$$

These differ from 4 by $\{3, 2, 4, 1, 0\} = \{0, 1, 2, 3, 4\}!$

Major index bijection

Corollary

There is a bijection

$$\Psi: \{\alpha \in \mathbb{Z}_{\geq 0}^n : (\alpha_1, \dots, \alpha_n) \leq (n-1, n-2, \dots, 1, 0)\} \rightarrow S_n$$

for which $\text{maj}(\Psi(\alpha)) = \alpha_1 + \dots + \alpha_n$.

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Use $\alpha = (2, 3, 0, 1, 0)$. Then:

	maj	Δ maj
1	0	0

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2.1	1	1

Major index bijection

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Use $\alpha = (2, 3, 0, 1, 0)$. Then:

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1	0	0
2.1	1	1
2.13	1	0

Major index bijection

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Use $\alpha = (2, 3, 0, 1, 0)$. Then:

	maj	Δ maj
1	0	0
2.1	1	1
2.13	1	0
2.14.3	4	3

Major index bijection

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Use $\alpha = (2, 3, 0, 1, 0)$. Then:

	maj	Δ maj
1	0	0
2.1	1	1
2.13	1	0
2.14.3	4	3
25.14.3	6	2

Hence $\Psi((2, 3, 0, 1, 0)) = 25143$.

Major index bijection

Corollary

*The bijection $\Psi \circ \Phi: S_n \rightarrow S_n$ sends inv to maj,
i.e. $\text{inv}(\pi) = \text{maj}((\Psi \circ \Phi)(\pi))$.*

Major index bijection

Corollary

The bijection $\Psi \circ \Phi: S_n \rightarrow S_n$ sends inv to maj, i.e. $\text{inv}(\pi) = \text{maj}((\Psi \circ \Phi)(\pi))$. Hence $\mathcal{X}_{\text{inv}} \sim \mathcal{X}_{\text{maj}}$ and \mathcal{X}_{maj} is asymptotically normal as $n \rightarrow \infty$.

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25.14.3	6	2

Hence $\Psi((2, 3, 0, 1, 0)) = 25143$. Since $\Phi(35142) = (2, 3, 0, 1, 0)$, we have

$$(\Psi \circ \Phi)(35142) = 25143$$

$$\text{inv}(35142) = 6 = \text{maj}(25143).$$

Inv and maj

Question (Svante Janson)

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Theorem (Baxter–Zeilberger)

inv and maj on S_n are *jointly independently asymptotically normally distributed* as $n \rightarrow \infty$. That is, for all $u, v \in \mathbb{R}$,

$$\lim_{n \rightarrow \infty} \mathbb{P}[\mathcal{X}_{\text{inv}}^* \leq u, \mathcal{X}_{\text{maj}}^* \leq v] = \frac{1}{2\pi} \int_{-\infty}^u \int_{-\infty}^v e^{-x^2/2} e^{-y^2/2} dy dx$$

where $\mathcal{X}^* := (\mathcal{X} - \mu_n)/\sigma_n$ with μ_n, σ_n from Feller's theorem.

Inv and maj

The Baxter–Zeilberger proof can be summarized as follows:

1. The *method of moments* says it suffices to show that for each fixed $(s, t) \in \mathbb{Z}_{\geq 0}^2$, the (s, t) -mixed moment $\mathbb{E}[(\mathcal{X}_{\text{inv}}^*)^s (\mathcal{X}_{\text{maj}}^*)^t]$ tend to the (s, t) -mixed moment of $\mathcal{N}(0, 1) \times \mathcal{N}(0, 1)$ as $n \rightarrow \infty$.

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2. Let $F_{n,i}(p, q) := \sum_{\pi_n=i} p^{\text{inv}(\pi)} q^{\text{maj}(\pi)}$. Derive a recurrence for $F_{n,i}(p, q)$ by considering the effect of removing the last letter.

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3. Use the recurrence and Taylor expansion to derive a recurrence for the *mixed factorial moments* $\mathbb{E}[(\mathcal{X}_{\text{inv}}^*)^{(s)} (\mathcal{X}_{\text{maj}}^*)^{(t)}]$.

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The details are involved and are perhaps best handled by a computer, which can easily compute all the relevant quantities using the recursions. The approach gives me no intuition for *why* the result should be true.

The \$300 question

“Referee Dan Romik believe[s] that we should mention, at this point, the ‘explicit’ formula of Roselle (mentioned by Knuth) in terms of a certain infinite double product for the q -exponential generating function of $\sum_{\pi \in S_n} p^{\text{inv}(\pi)} q^{\text{maj}(\pi)}$. Romik believes that this may lead to an alternative proof, that would even imply a stronger result (a local limit law). We strongly doubt this, and [Doron Zeilberger] is hereby offering \$300 for the first person to supply such a proof, whose length should not exceed the length of this article [13 pages].”

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Roselle's formula

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Theorem (Roselle)

We have

$$\sum_{n \geq 0} \frac{H_n(p, q) z^n}{(p)_n (q)_n} = \prod_{a, b \geq 0} \frac{1}{1 - p^a q^b z}$$

where $(p)_n := (1 - p)(1 - p^2) \cdots (1 - p^n)$.

A correction factor

If inv and maj on S_n were independent, we would have

$$\frac{H_n(p, q)}{n!} = \frac{[n]_p! [n]_q!}{n!^2}.$$

In this case, joint asymptotic normality would follow trivially from individual asymptotic normality.

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$$\frac{H_n(p, q)}{n!} = \frac{[n]_p! [n]_q!}{n!^2} F_n(p, q)$$

where

$$F_n(p, q) = \frac{n! \cdot \text{g.f. of size-}n \text{ multisets from } \mathbb{Z}_{\geq 0}^2}{\text{g.f. of size-}n \text{ lists from } \mathbb{Z}_{\geq 0}^2}.$$

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Intuitively, F_n is “1 to first order”. This explains “why” Baxter–Zeilberger's result holds and suggests an alternate proof.

Explicit correction factor

Theorem (S.)

There are constants $c_\mu \in \mathbb{Z}$ indexed by integer partitions μ such that

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$$F_n(p, q) = \sum_{d=0}^n [(1-p)(1-q)]^d \sum_{\substack{\mu \vdash n \\ \ell(\mu)=n-d}} \frac{c_\mu}{\prod_i [\mu_i]_p [\mu_i]_q}.$$

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Explicitly,

$$c_\mu = \sum_{\lambda \vdash n} \lambda! \sum_{\substack{\Lambda: \Pi(\lambda) \leq \Lambda \\ \text{type}(\Lambda) = \mu}} \mu(\Pi(\lambda), \Lambda).$$

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$$\frac{H_n(p, q)}{n!} = \frac{[n]_p! [n]_q!}{n!^2} F_n(p, q)$$

where

$$F_n(p, q) = \sum_{d=0}^n [(1-p)(1-q)]^d \sum_{\substack{\mu \vdash n \\ \ell(\mu)=n-d}} \frac{c_\mu}{\prod_i [\mu_i]_p [\mu_i]_q}.$$

Explicitly,

$$c_\mu = \sum_{\lambda \vdash n} \lambda! \sum_{\substack{\Lambda: \Pi(\lambda) \leq \Lambda \\ \text{type}(\Lambda) = \mu}} \mu(\Pi(\lambda), \Lambda).$$

The $d = 0$ contribution is 1. Hence, $H_n(1, q) = [n]_q!$.

Estimating the correction factor

Theorem (S.)

Uniformly on compact subsets of \mathbb{R}^2 , we have

$$F_n(e^{is/\sigma_n}, e^{it/\sigma_n}) \rightarrow 1 \quad \text{as} \quad n \rightarrow \infty$$

Estimating the correction factor

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The argument uses the explicit form of c_μ , the explicit form of the Möbius function on the lattice of set partitions, and some estimates to bound the $d > 0$ contributions to F_n .

Estimating the correction factor

Technical details: easy manipulations give, for $|s|, |t| \leq M$ and n large,

$$|F_n(e^{is/\sigma_n}, e^{it/\sigma_n}) - 1| \leq \sum_{d=1}^n \frac{|st|^d}{\sigma_n^{2d}} \sum_{\lambda \vdash n} \lambda! \sum_{\substack{\Lambda: \Pi(\lambda) \leq \Lambda \\ \#\Lambda = n-d}} |\mu(\Pi(\lambda), \Lambda)|.$$

Lemma

Suppose $\lambda \vdash n$ with $\ell(\lambda) = n - k$, and fix d . Then

$$\sum_{\substack{\Lambda: \Pi(\lambda) \leq \Lambda \\ \#\Lambda = n-d}} \mu(\Pi(\lambda), \Lambda) = (-1)^{d-k} \sum_{\substack{\Lambda \in P[n-k] \\ \#\Lambda = n-d}} \prod_{A \in \Lambda} (\#A - 1)!$$

and the terms on the left all have the same sign $(-1)^{d-k}$. The sums are empty unless $n \geq d \geq k \geq 0$.

Estimating the correction factor

Lemma

Let $\lambda \vdash n$ with $\ell(\lambda) = n - k$ and $n \geq d \geq k \geq 0$. Then

$$\sum_{\substack{\Lambda: \Pi(\lambda) \leq \Lambda \\ \#\Lambda = n-d}} |\mu(\Pi(\lambda), \Lambda)| \leq (n - k)^{2(d-k)}.$$

Lemma

For $n \geq d \geq k \geq 0$, we have

$$\sum_{\substack{\lambda \vdash n \\ \ell(\lambda) = n-k}} \lambda! \sum_{\substack{\Lambda: \Pi(\lambda) \leq \Lambda \\ \#\Lambda = n-d}} |\mu(\Pi(\lambda), \Lambda)| \leq (n - k)^{2d-k} (k + 1)!.$$

Estimating the correction factor

Lemma

For n sufficiently large, for all $0 \leq d \leq n$ we have

$$\sum_{\lambda \vdash n} \lambda! \sum_{\substack{\Lambda: \Pi(\lambda) \leq \Lambda \\ \#\Lambda = n-d}} |\mu(\Pi(\lambda), \Lambda)| \leq 3n^{2d}.$$

Putting it all together:

$$\begin{aligned} |F_n(e^{is/\sigma_n}, e^{it/\sigma_n}) - 1| &\leq 3 \sum_{d=1}^n \frac{(Mn)^{2d}}{\sigma_n^{2d}} \\ (Mn)^{2d} / \sigma_n^{2d} &\sim (36^2 M^2 / n)^d \\ \lim_{n \rightarrow \infty} \sum_{d=1}^n \frac{(Mn)^{2d}}{\sigma_n^{2d}} &= 0. \end{aligned}$$

Finishing up

Definition

The *characteristic function* of a real-valued random variable \mathcal{X} is

$$\begin{aligned}\phi_{\mathcal{X}}: \mathbb{R} &\rightarrow \mathbb{C} \\ \phi_{\mathcal{X}}(t) &:= \mathbb{E}[e^{i\mathcal{X}t}].\end{aligned}$$

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Theorem (*Lévy Continuity*)

$\mathcal{X}_1, \mathcal{X}_2, \dots$ *converges in distribution* to \mathcal{X} if and only if for all $t \in \mathbb{R}$,

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A similar result holds for \mathbb{R}^k -valued random variables.

Finishing up

The equation

$$\frac{H_n(p, q)}{n!} = \frac{[n]_p! [n]_q!}{n!^2} F_n(p, q)$$

can be reinterpreted as

$$\phi(\chi_{\text{inv}}^*, \chi_{\text{maj}}^*)(s, t) = \phi \chi_{\text{inv}}^*(s) \phi \chi_{\text{maj}}^*(t) F_n(e^{is/\sigma_n}, e^{it/\sigma_n}).$$

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For fixed s, t , using the theorem above and Feller's result gives

$$\lim_{n \rightarrow \infty} \phi(x_{\text{inv}}^*, x_{\text{maj}}^*)(s, t) = e^{-s^2/2} e^{-t^2/2} = \phi(\mathcal{N}(0,1), \mathcal{N}(0,1))(s, t).$$

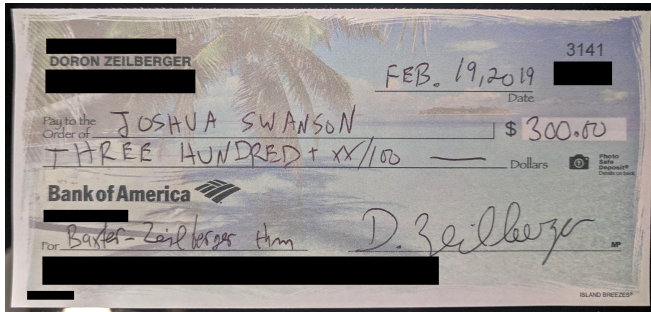
This completes the proof of the Baxter–Zeilberger theorem using Roselle's formula.

Done!

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Thanks!

${}_{\tau} \mathcal{H} \mathcal{A} \mathcal{N} \mathcal{K} \mathcal{S} !$

Extra slide: local limit theorem?

Romik's question was largely motivated by a desire to find a *local limit theorem*. Here, this would be a statement of the form

$$\mathbb{P}[\text{inv} = u, \text{maj} = v] = \frac{1}{2\pi\sigma_n} e^{-(u-\mu_n)^2/\sigma_n - (v-\mu_n)^2/\sigma_n} + O(f(n))$$

with an **explicit error bound** $f(n)$ where $\lim_{n \rightarrow \infty} f(n) = 0$.

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with an **explicit error bound** $f(n)$ where $\lim_{n \rightarrow \infty} f(n) = 0$.

The method of moments has no hope of proving such a result. A standard approach to local limit theorems is to use the **Cauchy integral formula** on the generating function, though such arguments are typically lengthy and technical. A local limit theorem in this context will be the subject of a future article.