

# Asymptotics of Mahonian statistics

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**Joshua P. Swanson**

University of California, San Diego

Based partly on joint work with

**Sara Billey** and **Matjaž Konvalinka**

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Slides: [http://www.math.ucsd.edu/~jswanson/talks/2019\\_SCDMS.pdf](http://www.math.ucsd.edu/~jswanson/talks/2019_SCDMS.pdf)

# Inversion number definition

## Definition

The *symmetric group* is

$$S_n := \{\text{bijections } \pi: \{1, 2, \dots, n\} \rightarrow \{1, 2, \dots, n\}\}.$$

The elements are *permutations*.

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For example,

$$\text{inv}(35142) = \#\{(1, 3), (1, 5), (2, 3), (2, 4), (2, 5), (4, 5)\} = 6.$$

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Zeilberger: this is the “**most important permutation statistic**”.

# Inversion number bijection

## Lemma (Classical)

*The map*

$$\Phi: S_n \rightarrow \{\alpha \in \mathbb{Z}_{\geq 0}^n : (\alpha_1, \dots, \alpha_n) \leq (n-1, n-2, \dots, 1, 0)\}$$
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	(2,5)		
(1,5)	(2,4)		
(1,3)	(2,3)		(4,5)

$$\Rightarrow \Phi(35142) = (2, 3, 0, 1, 0)$$

# Inversion number distribution

Clearly  $\text{inv}(\pi) = \alpha_1 + \cdots + \alpha_n$ . Hence:

## Corollary

- ▶  $\text{inv}$  on  $S_n$  is **symmetrically distributed** with mean  $((n-1) + (n-2) + \cdots + 0)/2 = \frac{1}{2}\binom{n}{2}$ .



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- ▶ The ordinary generating function of  $\text{inv}$  on  $S_n$  is

$$\begin{aligned}\sum_{\pi \in S_n} q^{\text{inv}(\pi)} &= \prod_{i=1}^n \sum_{\alpha_i=0}^{n-i} q^{\alpha_i} \\ &= (1 + q + \cdots + q^{n-1})(1 + q + \cdots + q^{n-2}) \cdots (1) \\ &=: [n]_q!\end{aligned}$$

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- ▶  $\mathcal{X}_{\text{inv}} \sim \mathcal{U}_{n-1} + \cdots + \mathcal{U}_1$  is the **sum of independent discrete uniform random variables**.

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Theorem (Feller '45; implicit earlier)

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$$\lim_{n \rightarrow \infty} \mathbb{P}[\mathcal{X}_{\text{inv}}^* \leq u] = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^u e^{-x^2/2} dx$$

where

$$\mathcal{X}_{\text{inv}}^* := \frac{\mathcal{X}_{\text{inv}} - \mu_n}{\sigma_n}$$

with

$$\mu_n = \frac{n(n-1)}{4}, \quad \sigma_n^2 = \frac{2n^3 + 3n^2 - 5n}{72}.$$

# Inversion number application

Theorem (Feller '45; implicit earlier)

As  $n \rightarrow \infty$ ,  $\mathcal{X}_{\text{inv}}$  is *asymptotically normal*.

Example (*Kendall's  $\tau$  test*)

Say some process generated distinct real numbers  $x_1, x_2, \dots, x_n$ , one per day. You want to know if the process is **independent of time**. Turn the data into a permutation  $\pi$  while preserving the relative order of data points and compute  $\text{inv}(\pi)$ . Since  $\mathcal{X}_{\text{inv}} \approx \mathcal{N}(\mu_n, \sigma_n)$ , independent data would have

$$|(\text{inv}(\pi) - \mu_n)/\sigma_n| \leq 3$$

$\approx 99.7\%$  of the time. So, **if this z-score is too big**, say larger than 3, the process is very likely time-dependent.

# Major index definition

Definition (MacMahon, early 1900's)

The *descent set* of  $\pi \in S_n$  is

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Zeilberger: this is the “**second most important permutation statistic**”.

# Major index bijection

## Lemma (Gupta, '78)

*For a given  $\pi \in S_{n-1}$ , let  $C_\pi \subset S_n$  be the  $n$  permutations obtained by inserting  $n$  into  $\pi$  in all possible ways. Then*

$$\{\text{maj}(\pi') - \text{maj}(\pi) : \pi' \in C_\pi\} = \{0, 1, \dots, n-1\}.$$

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## Corollary

*There is a bijection*

$$\Psi: \{\alpha \in \mathbb{Z}_{\geq 0}^n : (\alpha_1, \dots, \alpha_n) \leq (n-1, n-2, \dots, 1, 0)\} \rightarrow S_n$$

*for which  $\text{maj}(\Psi(\alpha)) = \alpha_1 + \dots + \alpha_n$ .*

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$\mathcal{X}_{\text{inv}} \sim \mathcal{X}_{\text{maj}}$  and  $\mathcal{X}_{\text{maj}}$  is asymptotically normal as  $n \rightarrow \infty$ .

# Major index bijection

## Example

Use  $\alpha = (2, 3, 0, 1, 0)$ . Then:

	maj	$\Delta$ maj
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25.14.3	6	2

Hence  $\Psi((2, 3, 0, 1, 0)) = 25143$ .

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25.14.3	6	2

Hence  $\Psi((2, 3, 0, 1, 0)) = 25143$ . Since  $\Phi(35142) = (2, 3, 0, 1, 0)$ , we have

$$(\Psi \circ \Phi)(35142) = 25143$$

$$\text{inv}(35142) = 6 = \text{maj}(25143).$$

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inv and maj on  $S_n$  are *jointly independently asymptotically normally distributed* as  $n \rightarrow \infty$ . That is, for all  $u, v \in \mathbb{R}$ ,

$$\lim_{n \rightarrow \infty} \mathbb{P}[\mathcal{X}_{\text{inv}}^* \leq u, \mathcal{X}_{\text{maj}}^* \leq v] = \frac{1}{2\pi} \int_{-\infty}^u \int_{-\infty}^v e^{-x^2/2} e^{-y^2/2} dy dx$$

where  $\mathcal{X}^* := (\mathcal{X} - \mu_n)/\sigma_n$  with  $\mu_n, \sigma_n$  from Feller's theorem.

## Inv and maj

The Baxter–Zeilberger proof can be summarized as follows:

1. The *method of moments* says it suffices to show that for each fixed  $(s, t) \in \mathbb{Z}_{\geq 0}^2$ , the  $(s, t)$ -mixed moment  $\mathbb{E}[(\mathcal{X}_{\text{inv}}^*)^s (\mathcal{X}_{\text{maj}}^*)^t]$  tend to the  $(s, t)$ -mixed moment of  $\mathcal{N}(0, 1) \times \mathcal{N}(0, 1)$  as  $n \rightarrow \infty$ .



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2. Let  $F_{n,i}(p, q) := \sum_{\pi_n=i} p^{\text{inv}(\pi)} q^{\text{maj}(\pi)}$ . Derive a recurrence for  $F_{n,i}(p, q)$  by considering the effect of removing the last letter.

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The details are involved and are perhaps best handled by a computer, which can easily compute all the relevant quantities using the recursions. The approach gives me no intuition for *why* the result should be true.

## The \$300 question

*“Referee Dan Romik believe[s] that we should mention, at this point, the ‘explicit’ formula of Roselle (mentioned by Knuth) in terms of a certain infinite double product for the  $q$ -exponential generating function of  $\sum_{\pi \in S_n} p^{\text{inv}(\pi)} q^{\text{maj}(\pi)}$ . Romik believes that this may lead to an alternative proof, that would even imply a stronger result (a local limit law). We strongly doubt this, and [Doron Zeilberger] is hereby offering \$300 for the first person to supply such a proof, whose length should not exceed the length of this article [13 pages].” (\*\*\*)*

# Roselle's formula

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## Theorem (Roselle)

We have

$$\sum_{n \geq 0} \frac{H_n(p, q) z^n}{(p)_n (q)_n} = \prod_{a, b \geq 0} \frac{1}{1 - p^a q^b z}$$

where  $(p)_n := (1 - p)(1 - p^2) \cdots (1 - p^n)$ .

## A correction factor

If inv and maj on  $S_n$  were independent, we would have

$$\frac{H_n(p, q)}{n!} = \frac{[n]_p! [n]_q!}{n!^2}.$$

In this case, joint asymptotic normality would follow trivially from individual asymptotic normality.



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$$\frac{H_n(p, q)}{n!} = \frac{[n]_p! [n]_q!}{n!^2} F_n(p, q)$$

where

$$F_n(p, q) = \frac{n! \cdot \text{g.f. of size-}n \text{ multisets from } \mathbb{Z}_{\geq 0}^2}{\text{g.f. of size-}n \text{ lists from } \mathbb{Z}_{\geq 0}^2}.$$

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Intuitively,  $F_n$  is “1 to first order”. This explains “why” Baxter–Zeilberger's result holds and suggests an alternate proof.

# Explicit correction factor

## Theorem (S.)

*There are constants  $c_\mu \in \mathbb{Z}$  indexed by integer partitions  $\mu$  such that*

$$\frac{H_n(p, q)}{n!} = \frac{[n]_p! [n]_q!}{n!^2} F_n(p, q)$$

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$$F_n(p, q) = \sum_{d=0}^n [(1-p)(1-q)]^d \sum_{\substack{\mu \vdash n \\ \ell(\mu)=n-d}} \frac{c_\mu}{\prod_i [\mu_i]_p [\mu_i]_q}.$$

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*Explicitly,*

$$c_\mu = \sum_{\lambda \vdash n} \lambda! \sum_{\substack{\Lambda: \Pi(\lambda) \leq \Lambda \\ \text{type}(\Lambda) = \mu}} \mu(\Pi(\lambda), \Lambda).$$

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Explicitly,

$$c_\mu = \sum_{\lambda \vdash n} \lambda! \sum_{\substack{\Lambda: \Pi(\lambda) \leq \Lambda \\ \text{type}(\Lambda) = \mu}} \mu(\Pi(\lambda), \Lambda).$$

The  $d = 0$  contribution is 1. Hence,  $H_n(1, q) = [n]_q!$ .

# Estimating the correction factor

## Theorem (S.)

*Uniformly on compact subsets of  $\mathbb{R}^2$ , we have*

$$F_n(e^{is/\sigma_n}, e^{it/\sigma_n}) \rightarrow 1 \quad \text{as} \quad n \rightarrow \infty$$

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The argument uses the explicit form of  $c_\mu$ , the explicit form of the Möbius function on the lattice of set partitions, and some estimates to bound the  $d > 0$  contributions to  $F_n$ .



# Estimating the correction factor

Technical details: easy manipulations give, for  $|s|, |t| \leq M$  and  $n$  large,

$$|F_n(e^{is/\sigma_n}, e^{it/\sigma_n}) - 1| \leq \sum_{d=1}^n \frac{|st|^d}{\sigma_n^{2d}} \sum_{\lambda \vdash n} \lambda! \sum_{\substack{\Lambda: \Pi(\lambda) \leq \Lambda \\ \#\Lambda = n-d}} |\mu(\Pi(\lambda), \Lambda)|.$$

## Lemma

Suppose  $\lambda \vdash n$  with  $\ell(\lambda) = n - k$ , and fix  $d$ . Then

$$\sum_{\substack{\Lambda: \Pi(\lambda) \leq \Lambda \\ \#\Lambda = n-d}} \mu(\Pi(\lambda), \Lambda) = (-1)^{d-k} \sum_{\substack{\Lambda \in P[n-k] \\ \#\Lambda = n-d}} \prod_{A \in \Lambda} (\#A - 1)!$$

and the terms on the left all have the same sign  $(-1)^{d-k}$ . The sums are empty unless  $n \geq d \geq k \geq 0$ .

# Estimating the correction factor

## Lemma

Let  $\lambda \vdash n$  with  $\ell(\lambda) = n - k$  and  $n \geq d \geq k \geq 0$ . Then

$$\sum_{\substack{\Lambda: \Pi(\lambda) \leq \Lambda \\ \#\Lambda = n-d}} |\mu(\Pi(\lambda), \Lambda)| \leq (n - k)^{2(d-k)}.$$

## Lemma

For  $n \geq d \geq k \geq 0$ , we have

$$\sum_{\substack{\lambda \vdash n \\ \ell(\lambda) = n-k}} \lambda! \sum_{\substack{\Lambda: \Pi(\lambda) \leq \Lambda \\ \#\Lambda = n-d}} |\mu(\Pi(\lambda), \Lambda)| \leq (n - k)^{2d-k} (k + 1)!.$$

# Estimating the correction factor

## Lemma

For  $n$  sufficiently large, for all  $0 \leq d \leq n$  we have

$$\sum_{\lambda \vdash n} \lambda! \sum_{\substack{\Lambda: \Pi(\lambda) \leq \Lambda \\ \#\Lambda = n-d}} |\mu(\Pi(\lambda), \Lambda)| \leq 3n^{2d}.$$

Putting it all together:

$$\begin{aligned} |F_n(e^{is/\sigma_n}, e^{it/\sigma_n}) - 1| &\leq 3 \sum_{d=1}^n \frac{(Mn)^{2d}}{\sigma_n^{2d}} \\ (Mn)^{2d} / \sigma_n^{2d} &\sim (36^2 M^2 / n)^d \\ \lim_{n \rightarrow \infty} \sum_{d=1}^n \frac{(Mn)^{2d}}{\sigma_n^{2d}} &= 0. \end{aligned}$$

# Finishing up

## Definition

The *characteristic function* of a real-valued random variable  $\mathcal{X}$  is

$$\begin{aligned}\phi_{\mathcal{X}}: \mathbb{R} &\rightarrow \mathbb{C} \\ \phi_{\mathcal{X}}(t) &:= \mathbb{E}[e^{i\mathcal{X}t}].\end{aligned}$$

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## Theorem (*Lévy Continuity*)

$\mathcal{X}_1, \mathcal{X}_2, \dots$  *converges in distribution* to  $\mathcal{X}$  if and only if for all  $t \in \mathbb{R}$ ,

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A similar result holds for  $\mathbb{R}^k$ -valued random variables.

## Finishing up

The equation

$$\frac{H_n(p, q)}{n!} = \frac{[n]_p! [n]_q!}{n!^2} F_n(p, q)$$

can be reinterpreted as

$$\phi(\chi_{\text{inv}}^*, \chi_{\text{maj}}^*)(s, t) = \phi \chi_{\text{inv}}^*(s) \phi \chi_{\text{maj}}^*(t) F_n(e^{is/\sigma_n}, e^{it/\sigma_n}).$$



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For fixed  $s, t$ , using the theorem above and Feller's result gives

$$\lim_{n \rightarrow \infty} \phi(x_{\text{inv}}^*, x_{\text{maj}}^*)(s, t) = e^{-s^2/2} e^{-t^2/2} = \phi(\mathcal{N}(0,1), \mathcal{N}(0,1))(s, t).$$

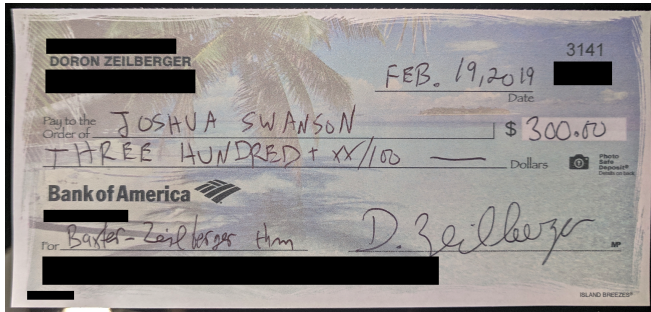
This completes the proof of the Baxter–Zeilberger theorem using Roselle's formula.

# Done!

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## Local limit theorem?

Romik's question was largely motivated by a desire to find a *local limit theorem*. Here, this would be a statement of the form

$$\mathbb{P}[\text{inv} = u, \text{maj} = v] = \frac{1}{2\pi\sigma_n} e^{-(u-\mu_n)^2/\sigma_n - (v-\mu_n)^2/\sigma_n} + O(f(n))$$

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The method of moments has no hope of proving such a result. A standard approach to local limit theorems is to use the **Cauchy integral formula** on the generating function, though such arguments are typically lengthy and technical. A local limit theorem in this context will be the subject of a future article.

# Variations

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## Question

What are the **possible normalized limit laws** for  $\text{maj}$  on **tableaux**?

# Partitions

## Definition

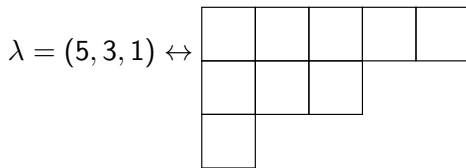
A *partition*  $\lambda$  of  $n$  is a sequence of positive integers  $\lambda_1 \geq \lambda_2 \geq \cdots$  such that  $\sum_i \lambda_i = n$ .



# Partitions

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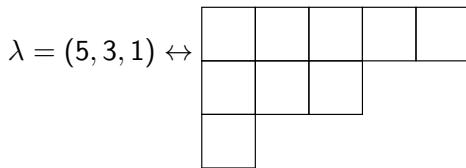
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## Theorem

(Young, early 1900's) *The complex inequivalent irreducible representations  $S^\lambda$  of  $S_n$  are **canonically** indexed by partitions of  $n$ .*

# Standard tableaux

## Definition

A *standard Young tableau* (*SYT*) of shape  $\lambda \vdash n$  is a filling of the cells of the Ferrers diagram of  $\lambda$  with  $1, 2, \dots, n$  which **increases along rows** and **decreases down columns**.

$$T = \begin{array}{|c|c|c|c|c|} \hline 1 & 3 & 6 & 7 & 9 \\ \hline 2 & 5 & 8 & & \\ \hline 4 & & & & \\ \hline \end{array} \in \text{SYT}(\lambda)$$

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Descent set:  $\{1, 3, 7\}$ .

## Definition

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$$\text{Des}(T) := \{1 \leq i < n : i+1 \text{ is in a lower row of } T \text{ than } i\}.$$

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The *major index* of  $T \in \text{SYT}(\lambda)$  is  $\text{maj}(T) := \sum_{i \in \text{Des}(T)} i$ .

## Standard tableaux

For  $\lambda = (5, 3, 1)$ ,

$$\begin{aligned} \sum_{T \in \text{SYT}(\lambda)} q^{\text{maj}(T)} &= q^5(q^{18} + 2q^{17} + 4q^{16} + 5q^{15} + 8q^{14} + 10q^{13} \\ &\quad + 13q^{12} + 14q^{11} + 16q^{10} + 16q^9 + 16q^8 + 14q^7 \\ &\quad + 13q^6 + 10q^5 + 8q^4 + 5q^3 + 4q^2 + 2q + 1). \end{aligned}$$

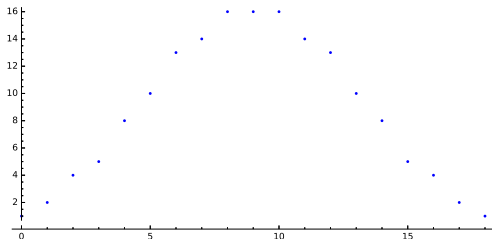
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The coefficients of  $q^{-5}f^{(5,3,1)}(q)$ :

(1, 2, 4, 5, 8, 10, 13, 14, 16, 16, 16, 14, 13, 10, 8, 5, 4, 2, 1)



## maj on SYT( $\lambda$ ) limit law classification

### Definition

Let  $\text{aft}(\lambda) := |\lambda| - \max\{\lambda_1, \lambda'_1\}$  and let

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- (i)  $\text{aft}(\lambda^{(n)}) \rightarrow \infty$ ; or
- (ii)  $|\lambda^{(n)}| \rightarrow \infty$  and  $\text{aft}(\lambda^{(n)}) \rightarrow M < \infty$ ; or
- (iii) the distribution of  $\mathcal{X}_{\lambda^{(n)}}^*[\text{maj}]$  is eventually constant.

The limit law is  $\mathcal{N}(0, 1)$  in case (i),  $\mathcal{IH}_M^*$  in case (ii), and *discrete* in case (iii).

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- ▶ Asymptotic normality is “obvious” when  $\sum_{j=1}^n j^d$  dominates, though for small  $\text{aft}(\lambda)$ , there is enormous cancellation resulting in degenerate cases with Irwin–Hall distributions.

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Thanks!

${}_{\tau} \mathcal{H} \mathcal{A} \mathcal{N} \mathcal{K} \mathcal{S} !$